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Spanish adaptation and validation of the Family Quality of Life Survey

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Abstract

Background Assessing the quality of life (QOL) for families that include a person with a disability have recently become a major emphasis in cross-cultural QOL studies. The present study examined the reliability and validity of the Family Quality of Life Survey (FQOL) on a Spanish sample.

Method and results The sample comprised 385 families who were administered the FQOL in Cali, Columbia. The FQOL showed adequate temporal stability (r = 0.68 on Importance and r = 0.78 on Satisfaction) and excellent internal consistency: Cronbach's alpha of 0.96 for Importance and 0.95 for Satisfaction. The confirmatory factor analysis yielded high fit indices, thus confirming that the factor structure of the FQOL as adapted for Spanish people fitted the five-factor model proposed by the survey's authors.

Conclusions The study provides a valid instrument for the research of the QOL of those families that have a child with a disability within Spanish-speaking community.

Keywords confirmatory factor analysis, family quality of life, intellectual disability, test adaptation

Introduction

Family assessment instruments have focused on variables such as stress and coping strategies, parenting, home environment, marital relationship or family needs. At best, the instruments to date have evaluated family satisfaction with the care services provided to the person with disability, but they have not analysed family quality of life (QOL) from a holistic perspective (Bailey *et al.* 1998; Browne & Bramster 1998; Ireys & Perry 1999; Rubio *et al.* 1999).

Over the last decade it has been suggested that the study of families with children with disabilities should modify its psychopathological approach and focus on the family assets to improve its potentials and capabilities (Turnbull *et al.* 1999; Wehman 2000; Turnbull & Turnbull 2001). The whole family is seen as a support unit (Schalock & Verdugo 2002) that collaborates with service providers and other families. The aim is not to 'treat' the family as the psychotherapeutic object but to meet their specific needs and expectations.

The Beach Center on Families and Disability at the University of Kansas has proposed a multidimensional model of QOL that includes domains and indicators focused on the person and the family (Park *et al.* 2003; Poston *et al.* 2003). Family QOL is measured by means of the Family Quality of Life Survey (FQOL), a tool that combines methodological pluralism (i.e. the integration of quantitative and quali-

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tative methods), solid theoretical rationale and suitable psychometric properties (Brown *et al.* 2003; Córdoba & Verdugo 2003; Park *et al.* 2003).

The QOL of the person with disabilities and that of the families are closely related and are influenced by personal and socio-cultural factors (Schalock & Verdugo 2002). A recent study on QOL (Schalock et al. 2005), which examined the results on importance and use of QOL indicators in three respondent groups (consumers, families and service providers/ professionals) and five geographical groupings (United States, Spain, Central and South America, Canada and Mainland China), showed that factors on importance and use scores are generally grouped into similar QOL domains, but there were significant differences in mean QOL importance and use scores for both the respondent and geographic groupings. This indicates that family QOL must be analysed independently and additionally to the study of person-centered QOL and, moreover, it must be carried out from a cross-cultural perspective.

Because of the recent development of the concept of family QOL in the Spanish language, there is a lack of measurement instruments with adequate psychometric characteristics. Therefore this study focused on the translation, adaptation and validation of the FQOL into the Spanish language according to the International Test Commission Guidelines for translating and adapting educational and psychological tests (Hambleton 1994).

Method

Participants

Participants were selected through a simple, random sampling procedure in several facilities of Cali (Colombia): the special-education institutions, the external clinic at the Child Hospital in Cali (Colombia), centres for attention and rehabilitation to children and adolescents with disabilities, and inclusive schools. The sample included more bi-parental than mono-parental families (64% vs. 32%), with a 4% presence of other types of family groupings. In regards to family lifespan, 57% of the families had children of school age or older (5–17 years of age), 23% had younger children (less than 5 years of age), and 20% had adult daughters and sons (over 18 years old). In regards to socio-economic status, 56% of the families had a low status, 34% had a medium status, and in 10% of the cases the status was high. The age of the family member with a disability was 3-5 in 12% of the families, 6-11 in 62% of the families, and 12– 17 in 26%. Finally, 58% of the persons with a disability participating in the study were male and 42% were female.

The inquiry consisted of 385 families of children and teenagers with a disability (aged 3–17) from the city of Cali, Colombia. Total 29.9% were persons with an intellectual disability, 23.9% with a sensory disability, 15.6% with a physical disability, 13% with a learning disability, 10% with attention deficit disorder and 6.8% with multiple disabilities. As for the respondents, 71.93% were mothers, 12.45% fathers, 3.11% brothers and sisters and 12.2% other members of the family.

Instrument

The Family Quality of Life Survey (Beach Center 2001; Park *et al.* 2003; Poston *et al.* 2003) was used to determine the satisfaction the family feels about different indicators of the QOL in contrast with the importance they attach to those indicators. The scale consists of 41 items grouped in five different factors: Family interaction, Parenting, Health and safety, Family resources and Support for persons with disabilities.

Procedure

The study was developed in four stages: (I) back translating and agreement carried out by eight translators; (2) evaluating the content of the items and its adaptation to the five factors, carried out by IO expert judges on the subject who classified each of the items by its category and intensity; (3) completing the socio-demographic survey and interviewers training; and (4) administering the scale by five psychologists during I6 months. The individualized administration of the scale took 45 min in the course of a direct interview. Informed consent was obtained from each participant and confidentiality of the data was guaranteed.

Analysis

Two studies were carried out to determine the reliability of the instrument. First, the test-retest proce-

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dure was used to evaluate the stability of measures with a subsample comprising 35 subjects from the overall sample, and using a time interval of 3– 4 weeks. Second, Cronbach's alpha indices were established on the total sample in order to estimate the internal consistency of the instrument.

Two studies were also conducted to test the validity of the survey: analysis of content and dimensional structure. Content validity was determined on the basis of expert opinion and comparison with the theoretical framework. The dimensional structure proposed by the authors was tested by means of a confirmatory factor analysis (CFA) using the Lisrel 8.3 program (Jöreskog & Sörbom 1996a). As the item scores are ordinal, polychoric correlations were first calculated using Prelis 2 (Jöreskog & Sörbom 1996b).

Results

Reliability

The results of the reliability analysis are shown in Table I. All the coefficients were significant (P < 0.0I) and most of them were greater than 0.60. The overall scale coefficients, for both Importance (0.68) and Satisfaction (0.78), enable us to conclude that the scale has adequate temporal stability.

The results for the analysis of the instrument's internal consistency are shown in the last two columns of Table I. All the alpha values are very high, with the total scale value illustrating that the scale has excellent internal consistency for both Importance (0.96) and Satisfaction (0.95). The corrected alpha coefficients were also calculated by eliminating each one of the scale items in turn; none of these procedures yielded values higher than the alpha coefficients shown in Table I. This indicates that the items have good discriminative power and that none of them can be considered eliminable.

Validity

There was very high agreement among the expert judges in terms of categorization: 70% to 90% agreement for most of the items. However, discrepancies were found on items 1, 3, 14, 37 and 39 (between 30 and 50% agreement). Given the judges' observations and the theoretical basis on which the scale was constructed, the text was altered substantially in these five items although they were maintained in the categories defined by the authors so as not to affect the initial structure of the instrument. In terms of the degree to which each item measured the corresponding category, the expert judges assigned a high rating to the 41 items (mean between 4 and 5 for 38 of the items); the items with the lowest rating were numbers 17 (3.25), 10 (3.50) and 32 (3.60), although these means should nonetheless be considered as high.

In the analysis of the dimensional structure using confirmatory methodology, the structure proposed by the authors was specified and tested to verify whether the data obtained in the present study fitted the original model. All the factor loadings of the 41 items were greater than 0.50, indicating their relevance to the measurement of the respective category.

| Table I (| Test-retest | coefficients | and | alfa | coefficients |
|-----------|-------------|--------------|-----|------|--------------|
|-----------|-------------|--------------|-----|------|--------------|

| | Test-retest | | Alfa coefficient | | |
|---------------------------------------|-------------|--------------|------------------|--------------|--|
| Subscales | Importance | Satisfaction | Importance | Satisfaction | |
| Family interaction | 0.485 | 0.779 | 0.838 | 0.873 | |
| Parenting | 0.610 | 0.464 | 0.857 | 0.855 | |
| Health and safety | 0.681 | 0.706 | 0.890 | 0.855 | |
| Family resources | 0.681 | 0.747 | 0.890 | 0.874 | |
| Support for persons with disabilities | 0.687 | 0.769 | 0.897 | 0.837 | |
| Total score | 0.684 | 0.785 | 0.959 | 0.946 | |

P < 0.01.

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Table 2 CFA adjustment rates for Importance and Satisfaction

| Survey | χ² (P) | RMR | GFI | NFI | RFI |
|--------------|-----------------|-------|------|------|------|
| Importance | 748.22 (0.70) | 0.048 | 0.99 | 0.99 | 0.99 |
| Satisfaction | 96.9 (< 0.00) | 0.060 | 0.98 | 0.97 | 0.97 |

CFA, confirmatory factor analysis; GFI, goodness-of-fit index; NFI, normed fit index; RFI, relative fit index; RMR, root mean square residual.

For both Importance and Satisfaction, the Support for persons with disabilities factor yielded the highest loadings, while the Family interaction factor showed a lower mean loading than the other categories.

In terms of the model's fit to the data, Table 2 shows the respective values of the χ^2 statistic, the root mean square residual (RMR), the goodness-of-fit index (GFI), the normed fit index (NFI) and the relative fit index (RFI). A good fit is assumed if χ^2 is not significant, the RMR is close to 0, and the GFI, NFI and RFI are close to 1. In the case of the Importance scale, all the indices reached optimum fit values. For the Satisfaction scale the χ^2 statistic was significant, while the other indices yielded very similar values to those obtained for the importance scale. Given the sensitivity of the χ^2 statistic to deviations from normality Jöreskog & Sörbom (1996a) recommend placing greater value on the other indices. In sum, it can be concluded that the model's fit to the data yields good values that enable us to accept the specification of the model for both Importance and Satisfaction.

Discussion

Research on the adaptation to the Spanish language of the assessment instrument for FQOL allows us to draw the conclusion that the model suggested by its authors (Park *et al.* 2003; Poston *et al.* 2003) has been successfully validated, keeping the original factorial structure and items. Quality of life has different meanings for every culture and person, and it differs by context, place and time (Matsumoto 2000). This confirmation of the factor structure in a different country with a different language and culture supports the etic/universal property of the family QOL construct (Keith & Schalock 2000; Schalock & Verdugo 2002; Skevington 2002; Schalock *et al.* 2005).

Family QOL instruments, such as the one validated in this study, are increasingly based on a better understanding of the core domains constituting a family's QOL and the core indicators that can be measured and used for multiple purposes. Because of the crosscultural validation of these domains and indicators, the goal of an enhanced family's QOL can serve as the basis for the development of policies, plans, projects and programmes orientated to favour the enhancement of persons with ID and their families.

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